

The exact asymptotics for hitting probability of a remote orthant by a multivariate Lévy process: the Cramér case

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Abstract For a multivariate Lévy process satisfying the Cramér moment condition and having a drift vector with at least one negative component, we derive the exact asymptotics of the probability of ever hitting the positive orthant that is being translated to infinity along a fixed vector with positive components. This problem is motivated by the multivariate ruin problem introduced in F. Avram et al. (2008) in the two-dimensional case. Our solution relies on the analysis from Y. Pan and K. Borovkov (2017) for multivariate random walks and an appropriate time discretization.

1 Introduction

In this note we consider the following large deviation problem for continuous time processes with independent increments that was motivated by the multivariate simultaneous ruin problem introduced in [1]. Let $\{\mathbf{X}(t)\}_{t \geq 0}$ be a d -dimensional ($d \geq 2$) right-continuous Lévy process with $\mathbf{X}(0) = 0$. One is interested in finding the precise asymptotics for the hitting probability of the orthant sG as $s \rightarrow \infty$, where

$$G := \mathbf{g} + Q^+$$

for some fixed

$$\mathbf{g} \in Q^+, \quad Q^+ := \{\mathbf{x} = (x_1, \dots, x_d) \in \mathbb{R}^d : x_j > 0, 1 \leq j \leq d\}.$$

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Clearly, $sG = s\mathbf{g} + Q^+$, which is just the positive orthant translated by $s\mathbf{g}$.

We solve this problem under appropriate Cramér moment assumptions and further conditions on the process \mathbf{X} and vertex \mathbf{g} that, roughly speaking, ensure that the “most likely place” for \mathbf{X} to hit sG when s is large is in vicinity of the “corner point” $s\mathbf{g}$. More specifically, we show that the precise asymptotics of the hitting probability of sG are given by the following expression: letting

$$\tau(V) := \inf\{t \geq 0 : \mathbf{X}(t) \in V\}$$

be the first hitting time of the set $V \subset \mathbb{R}^d$ by the process \mathbf{X} , one has

$$\mathbf{P}(\tau(sG) < \infty) = A_0 s^{-(d-1)/2} e^{-sD(G)} (1 + o(1)) \quad \text{as } s \rightarrow \infty, \quad (1)$$

where the “adjustment coefficient” $D(G)$ is the value of the second rate function (see (9) below) for the distribution of $\mathbf{X}(1)$ on the set G and the constant $A_0 \in (0, \infty)$ can be computed explicitly.

The asymptotics (1) extend a number of known results. The main body of literature on the topic of precise asymptotics for boundary crossing large deviation probabilities in the multivariate case concerns the random walk theory, see [3, 4, 5] and references therein for an overview of the relevant results. The crude logarithmic asymptotics in the multivariate case was also derived independently in [6].

The entrance probability to a remote set for Lévy processes was analyzed later, usually under some specific assumptions on the structure of these processes. For example, paper [1] dealt with the two-dimensional reserve process of the form

$$\mathbf{X}(t) = (X_1(t), X_2(t)) = (c_1, c_2) \sum_{i=1}^{N(t)} C_i - (p_1, p_2)t, \quad t \geq 0, \quad (2)$$

where $c_i, p_i > 0$, $i = 1, 2$, are constants, $\{C_i\}_{i \geq 1}$ is a sequence of i.i.d. claim sizes, and $N(t)$ is an independent of the claim sizes Poisson process. That model admits the following interpretation: the components of the process $s\mathbf{g} - \mathbf{X}$ describe the dynamics of the reserves of two insurance companies that start with the initial reserves sg_1 and sg_2 , respectively, and then divide between them both claims and premia in some pre-specified proportions. In that case, $\mathbf{P}(\tau(sG) < \infty)$ corresponds to the simultaneous ruin probability of the two companies. The main result of the present paper generalizes the assertion of Theorem 5 of [1] to the case of general Lévy processes. One may also wish to mention here the relevant papers [2, 7].

2 The Main Result

To state the main result, we will need some notations. For brevity, denote by $\boldsymbol{\xi}$ a random vector such that

$$\boldsymbol{\xi} \stackrel{d}{=} \mathbf{X}(1). \quad (3)$$

Our first condition on \mathbf{X} is stated as follows.

[C₁] *The distribution of ξ is non-lattice and there is no hyperplane $H = \{\mathbf{x} : \langle \mathbf{a}, \mathbf{x} \rangle = c\} \subset \mathbb{R}^d$ such that $\mathbf{P}(\xi \in H) = 1$.*

That condition can clearly be re-stated in terms of the covariance matrix of the Brownian component and spectral measure of \mathbf{X} , although such re-statement will not make it more compact nor transparent.

Next denote by

$$K(\boldsymbol{\lambda}) := \ln \mathbf{E} e^{\langle \boldsymbol{\lambda}, \xi \rangle}, \quad \boldsymbol{\lambda} \in \mathbb{R}^d, \quad (4)$$

the cumulant function of ξ and let

$$\Theta_\psi := \{\boldsymbol{\lambda} \in \mathbb{R}^d : K(\boldsymbol{\lambda}) < \infty\}$$

be the set on which the moment generating function of ξ is finite. We will need the following Cramér moment condition on \mathbf{X} :

[C₂] Θ_ψ contains a non-empty open set.

The first rate function $\Lambda(\boldsymbol{\alpha})$ for the random vector ξ is defined as the Legendre transform of the cumulant function K :

$$\Lambda(\boldsymbol{\alpha}) := \sup_{\boldsymbol{\lambda} \in \Theta_\psi} (\langle \boldsymbol{\alpha}, \boldsymbol{\lambda} \rangle - K(\boldsymbol{\lambda})), \quad \boldsymbol{\alpha} \in \mathbb{R}^d. \quad (5)$$

The probabilistic interpretation of the first rate function is given by the following relation (see e.g. [5]): for any $\boldsymbol{\alpha} \in \mathbb{R}^d$,

$$\Lambda(\boldsymbol{\alpha}) = - \lim_{\varepsilon \rightarrow 0} \lim_{n \rightarrow \infty} \frac{1}{n} \ln \mathbf{P} \left(\frac{\mathbf{X}(n)}{n} \in U_\varepsilon(\boldsymbol{\alpha}) \right), \quad (6)$$

where $U_\varepsilon(\boldsymbol{\alpha})$ is the ε -neighborhood of $\boldsymbol{\alpha}$. Accordingly, for a set $B \subset \mathbb{R}^d$, any point $\boldsymbol{\alpha} \in B$ such that

$$\Lambda(\boldsymbol{\alpha}) = \Lambda(B) := \inf_{\mathbf{v} \in B} \Lambda(\mathbf{v}) \quad (7)$$

is called a most probable point (MPP) of the set B (cf. relation (11) in [8]). If such a point $\boldsymbol{\alpha}$ is unique for a given set B , we denote it by

$$\boldsymbol{\alpha}[B] := \arg \min_{\mathbf{v} \in B} \Lambda(\mathbf{v}). \quad (8)$$

Now recall the definition of the second rate function D that was introduced and studied in [3]: letting $D_u(\mathbf{v}) := u\Lambda(\mathbf{v}/u)$ for $\mathbf{v} \in \mathbb{R}^d$, one sets

$$D(\mathbf{v}) := \inf_{u > 0} D_u(\mathbf{v}), \quad \mathbf{v} \in \mathbb{R}^d, \quad D(B) := \inf_{\mathbf{v} \in B} D(\mathbf{v}), \quad B \subset \mathbb{R}^d \quad (9)$$

(see also [8]). Further, we put

$$r_B := \arg \min_{r > 0} D_{1/r}(B). \quad (10)$$

Recall the probabilistic meaning of the function D and the value r_B . While the first rate function Λ specifies the main term in the asymptotics of the probabilities for the random walk values $\mathbf{X}(n)$ to be inside “remote sets” (roughly speaking, $\Lambda(B)$ equals the RHS of (6) with the neighbourhood of $\boldsymbol{\alpha}$ in it replaced with B), the second rate function D does that for the probabilities of *ever hitting* “remote sets” by the whole random walk trajectory $\{\mathbf{X}(n)\}_{n \geq 0}$, the meaning of r_B being that $1/r_B$ gives (after appropriate scaling) the “most probable time” for the walk to hit the respective remote set. For more detail, we refer the interested reader to [3, 8].

Define the Cramér range Ω_Λ for $\boldsymbol{\xi}$ as follows:

$$\Omega_\Lambda := \{ \boldsymbol{\alpha} = \text{grad}K(\boldsymbol{\lambda}) : \boldsymbol{\lambda} \in \text{int}(\Theta_\Psi) \},$$

where the cumulant function $K(\boldsymbol{\lambda})$ of $\boldsymbol{\xi}$ was defined in (4) and $\text{int}(B)$ stands for the interior of the set B . In words, the set Ω_Λ consists of all the vectors that can be obtained as the expectations of the Cramér transforms of the law of $\boldsymbol{\xi}$, i.e. the distributions of the form $e^{(\boldsymbol{\lambda}, \mathbf{x}) - K(\boldsymbol{\lambda})} \mathbf{P}(\boldsymbol{\xi} \in d\mathbf{x})$, for parameter values $\boldsymbol{\lambda} \in \text{int}(\Theta_\Psi)$.

For $\boldsymbol{\alpha} \in \mathbb{R}^d$, denote by $\boldsymbol{\lambda}(\boldsymbol{\alpha})$ the vector $\boldsymbol{\lambda}$ at which the upper bound in (5) is attained (when such a vector exists, in which case it is always unique):

$$\Lambda(\boldsymbol{\alpha}) = \langle \boldsymbol{\alpha}, \boldsymbol{\lambda}(\boldsymbol{\alpha}) \rangle - K(\boldsymbol{\lambda}(\boldsymbol{\alpha})).$$

For $r > 0$, assuming that $\boldsymbol{\alpha}[rG] \in \Omega_\Lambda$, introduce the vector

$$\mathbf{N}(r) := \text{grad}\Lambda(\boldsymbol{\alpha})|_{\boldsymbol{\alpha}=\boldsymbol{\alpha}[rG]} = \boldsymbol{\lambda}(\boldsymbol{\alpha}[rG]), \quad (11)$$

which is a normal to the level surface of Λ at the point $\boldsymbol{\alpha}[rG]$ (see e.g. (22) in [8]).

The last condition that we will need to state our main result depends on the parameter $r > 0$ and is formulated as follows:

[C₃(r)] *One has*

$$\Lambda(rG) = \Lambda(r\mathbf{g}), \quad r\mathbf{g} \in \Omega_\Lambda, \quad \mathbf{N}(r) \in Q^+, \quad \langle \mathbf{E}\boldsymbol{\xi}, \mathbf{N}(r) \rangle < 0.$$

The first part of condition [C₃(r)] means that the vertex $r\mathbf{g}$ is an MPP for the set rG . Note that under the second part of the condition, this MPP $r\mathbf{g}$ for rG is unique (e.g., by Lemma 1 in [8]). Since $\mathbf{N}(r)$ always belongs to the closure of Q^+ , the third part of condition [C₃(r)] just excludes the case when the normal $\mathbf{N}(r)$ to the level surface of Λ at the point $r\mathbf{g}$ belongs to the boundary of the set rG .

Theorem 1. *Let conditions [C₁], [C₂] and [C₃(r_G)] be met. Then the asymptotic relation (1) holds true, where $D(G)$ is the value of the second rate function (9) on G and the constant $A_0 \in (0, \infty)$ can be computed explicitly.*

The value of the constant $A_0 \in (0, \infty)$ is given by the limit as $\delta \rightarrow 0$ of the expressions given by formula (68) in [8] for the distribution of $\boldsymbol{\xi} \stackrel{d}{=} \mathbf{X}(\delta)$. When proving the theorem below, we demonstrate that that limit does exist and is finite and positive.

Proof. For a $\delta > 0$, consider the embedded random walk $\{\mathbf{X}(n\delta)\}_{n \in \mathbb{N}}$ and, for a set $V \subset \mathbb{R}^d$, denote the first time that random walk hits that set V by

$$\eta_\delta(V) := \inf\{n \in \mathbb{N} : \mathbf{X}(n\delta) \in V\}.$$

First observe that, on the one hand, for any $\delta > 0$, one clearly has

$$\mathbf{P}(\tau(sG) < \infty) \geq \mathbf{P}(\eta_\delta(sG) < \infty). \quad (12)$$

On the other hand, assuming without loss of generality that $\min_{1 \leq j \leq d} g_j \geq 1$ and setting $I(s) := (\tau(sG), \tau(sG) + \delta] \subset \mathbb{R}$ on the event $\{\tau(sG) < \infty\}$, we have, for any $\varepsilon > 0$,

$$\begin{aligned} \mathbf{P}(\eta_\delta((s - \varepsilon)G) < \infty) &\geq \mathbf{P}\left(\tau(sG) < \infty, \sup_{t \in I(s)} \|\mathbf{X}(t) - \mathbf{X}(\tau(sG))\| \leq \varepsilon\right) \\ &= \mathbf{P}(\tau(sG) < \infty) \mathbf{P}\left(\sup_{t \in I(s)} \|\mathbf{X}(t) - \mathbf{X}(\tau(sG))\| \leq \varepsilon\right) \\ &= \mathbf{P}(\tau(sG) < \infty) \mathbf{P}\left(\sup_{t \in (0, \delta]} \|\mathbf{X}(t)\| \leq \varepsilon\right), \end{aligned} \quad (13)$$

where the last two relations follow from the strong Markov property and homogeneity of \mathbf{X} .

Now take an arbitrary small $\varepsilon > 0$. As the process \mathbf{X} is right-continuous, there exists a $\delta(\varepsilon) > 0$ such that

$$\mathbf{P}\left(\sup_{t \in (0, \delta(\varepsilon)]} \|\mathbf{X}(t)\| \leq \varepsilon\right) > (1 + \varepsilon)^{-1},$$

which, together with (13), yields the inequality

$$\mathbf{P}(\tau(sG) < \infty) \leq (1 + \varepsilon) \mathbf{P}(\eta_\delta((s - \varepsilon)G) < \infty). \quad (14)$$

The precise asymptotics of the probability on the RHS of (12) were obtained in [8]. It is given in terms of the second rate function $D^{[\delta]}$ for the distribution of the jumps $\mathbf{X}(n\delta) - \mathbf{X}((n-1)\delta) \stackrel{d}{=} \mathbf{X}(\delta)$ in the random walk $\{\mathbf{X}(n\delta)\}_{n \geq 0}$. Recalling the well-known fact that the cumulant of $\mathbf{X}(\delta)$ is given by δK , we see that the first rate function $\Lambda^{[\delta]}$ for $\mathbf{X}(\delta)$ equals

$$\begin{aligned} \Lambda^{[\delta]}(\boldsymbol{\alpha}) &= \sup_{\boldsymbol{\lambda} \in \Theta_\psi} (\langle \boldsymbol{\alpha}, \boldsymbol{\lambda} \rangle - \delta K(\boldsymbol{\lambda})) \\ &= \delta \sup_{\boldsymbol{\lambda} \in \Theta_\psi} (\langle \boldsymbol{\alpha}/\delta, \boldsymbol{\lambda} \rangle - K(\boldsymbol{\lambda})) = \delta \Lambda(\boldsymbol{\alpha}/\delta), \quad \boldsymbol{\alpha} \in \mathbb{R}^d \end{aligned}$$

(cf. (5)). Therefore the second rate function (see (9)) $D^{[\delta]}$ for $\mathbf{X}(\delta)$ is

$$D^{[\delta]}(\mathbf{v}) := \inf_{u > 0} u \Lambda^{[\delta]}(\mathbf{v}/u) = \inf_{u > 0} (u\delta) \Lambda(\boldsymbol{\alpha}/(u\delta)) = D(\mathbf{v}), \quad \mathbf{v} \in \mathbb{R}^d.$$

That is, the second rate function for the random walk $\{\mathbf{X}(n\delta)\}$ is the same for all $\delta > 0$, which makes perfect sense as one would expect the same asymptotics for the probabilities $\mathbf{P}(\eta_\delta(sG) < \infty)$ for different δ . Hence the respective value $r_G^{[\delta]}$ (see (10)) can easily be seen to be given by δr_G .

Therefore, applying Theorem 1 in [8] to the random walk $\{\mathbf{X}(n\delta)\}$ and using notation $A^{[\delta]}$ for the constant A appearing in that theorem for the said random walk, we conclude that, for any $\delta \in (0, \delta(\varepsilon)]$, as $s \rightarrow \infty$,

$$\mathbf{P}(\eta_\delta(sG) < \infty) = A^{[\delta]} s^{-(d-1)/2} e^{-sD(G)} (1 + o(1)), \quad (15)$$

and likewise

$$\mathbf{P}(\eta_\delta((s-\varepsilon)G) < \infty) = A^{[\delta]} (s-\varepsilon)^{-(d-1)/2} e^{-(s-\varepsilon)D(G)} (1 + o(1)). \quad (16)$$

Now from (12) and (14) we see that, as $s \rightarrow \infty$,

$$A^{[\delta]} (1 + o(1)) \leq R(s) := \frac{\mathbf{P}(\tau(sG) < \infty)}{s^{-(d-1)/2} e^{-sD(G)}} \leq \frac{A^{[\delta]} (1 + \varepsilon) e^{\varepsilon D(G)}}{(1 - \varepsilon/s)^{(d-1)/2}} (1 + o(1)).$$

Therefore, setting $\underline{R} := \liminf_{s \rightarrow \infty} R(s)$, $\bar{R} := \limsup_{s \rightarrow \infty} R(s)$, we have

$$A^{[\delta]} \leq \underline{R} \leq \bar{R} \leq (1 + \varepsilon) e^{\varepsilon D(G)} A^{[\delta]}$$

for any $\delta \in (0, \delta(\varepsilon)]$, and hence

$$\limsup_{\delta \rightarrow 0} A^{[\delta]} \leq \underline{R} \leq \bar{R} \leq (1 + \varepsilon) e^{\varepsilon D(G)} \liminf_{\delta \rightarrow 0} A^{[\delta]}.$$

As $\varepsilon > 0$ is arbitrary small, we conclude that there exists $\lim_{\delta \rightarrow 0} A^{[\delta]} =: A_0 \in (0, \infty)$. Therefore there also exists

$$\lim_{s \rightarrow \infty} R(s) = A_0.$$

The theorem is proved. \square

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